

Q. Let $X_1, X_2, \dots, X_n \stackrel{iid}{\sim} N(\mu, \sigma^2)$. Let $s^2 = \frac{1}{n-1} \sum (x_i - \bar{x})^2$

Show that s^2 is a consistent estimator of σ^2 .

check: $E(s^2) = \sigma^2$ and $Var(s^2) \rightarrow 0$ as $n \rightarrow \infty$.

$$s^2 = \frac{1}{n-1} \sum (x_i - \bar{x})^2$$

$$E(s^2) = E \left[\frac{1}{n-1} \sum (x_i - \bar{x})^2 \right] = \sigma^2$$

$$Var(s^2) = Var \left[\frac{1}{n-1} \sum (x_i - \bar{x})^2 \right]$$

$$= \frac{1}{(n-1)^2} Var \left[\sum (x_i - \bar{x})^2 \right]$$

$$= \frac{1}{(n-1)^2} Var \left[\sigma^2 \sum \left(\frac{x_i - \bar{x}}{\sigma} \right)^2 \right]$$

$$= \frac{\sigma^4}{(n-1)^2} Var \left[\sum \left(\frac{x_i - \bar{x}}{\sigma} \right)^2 \right]$$

$$= \frac{\sigma^4}{(n-1)^2} \cdot 2(n-1) = \frac{2\sigma^4}{(n-1)}$$

$X_1, X_2, \dots, X_n \sim N(\mu, \sigma^2)$

$\left(\frac{x_i - \mu}{\sigma} \right) \sim N(0, 1)$

$\sum \left(\frac{x_i - \mu}{\sigma} \right)^2 \sim \chi^2_{(n)}$

$\hookrightarrow E[\cdot] = n$

$Var[\cdot] = 2n$

$\sum \left(\frac{x_i - \bar{x}}{\sigma} \right)^2 \sim \chi^2_{(n-1)}$

$\hookrightarrow E[\cdot] = (n-1)$

$Var[\cdot] = 2(n-1)$

As $n \rightarrow \infty$, $Var(s^2) \rightarrow 0$. $\therefore s^2$ is consistent for σ^2 .

Q. Let X_1, X_2, \dots, X_n be a r.v.s from $U[0, \theta]$. Define $T = \left(\prod x_i \right)^{1/n}$. Show that T is a consistent estimator of (θ/e) .

pdf of $U[0, \theta]$ $f(x) = \frac{1}{\theta}, 0 \leq x \leq \theta$

[Note: $X \sim U[a, b]$ $E(X) = \frac{a+b}{2}$, $Var(X) = \frac{(a-b)^2}{12}$

$f(x) = \frac{1}{b-a}$, $a \leq x \leq b$

$E(X) = \int_a^b x f(x) dx$
 $E(X^2) = \int_a^b x^2 f(x) dx$
 $Var(X) = E(X^2) - [E(X)]^2$

$T = (x_1 \cdot x_2 \cdot \dots \cdot x_n)^{\frac{1}{n}}$

$E(T) = E[(x_1 \cdot x_2 \cdot \dots \cdot x_n)^{\frac{1}{n}}]$

$= E[x_1^{\frac{1}{n}} x_2^{\frac{1}{n}} \dots x_n^{\frac{1}{n}}]$

$= E(x_1^{\frac{1}{n}}) \cdot E(x_2^{\frac{1}{n}}) \cdot \dots \cdot E(x_n^{\frac{1}{n}})$

$= \prod E(x_i^{\frac{1}{n}})$

$= \prod \frac{\theta^{\frac{1}{n}}}{(1+\frac{1}{n})} = \left(\frac{\theta^{\frac{1}{n}}}{(1+\frac{1}{n})}\right)^n = \frac{\theta}{(1+\frac{1}{n})^n}$

$E(x_i^{\frac{1}{n}}) = \int_a^b x_i^{\frac{1}{n}} f(x) dx$

$= \int_0^{\theta} x^{\frac{1}{n}} \cdot \frac{1}{\theta} dx$
 $= \frac{1}{\theta} \int_0^{\theta} x^{\frac{1}{n}} dx$

$= \frac{1}{\theta} \left[\frac{x^{\frac{1}{n}+1}}{\frac{1}{n}+1} \right]_0^{\theta}$

$= \frac{1}{\theta} \left[\frac{\theta^{\frac{1}{n}+1}}{\frac{1}{n}+1} \right]$

$= \frac{\theta^{\frac{1}{n}}}{(1+\frac{1}{n})}$

$\lim_{n \rightarrow \infty} E(T) = \lim_{n \rightarrow \infty} \frac{\theta}{(1+\frac{1}{n})^n} = \left(\frac{\theta}{e}\right)$

$T = \left(\prod_{i=1}^n x_i\right)^{\frac{1}{n}}$, $\lim_{n \rightarrow \infty} E(T) \rightarrow \left(\frac{\theta}{e}\right)$

$E(T^2) = E\left[\left(\prod_{i=1}^n x_i\right)^{\frac{2}{n}}\right]$

$= \prod_{i=1}^n E(x_i^{\frac{2}{n}})$

$= \prod \frac{\theta^{\frac{2}{n}}}{(1+\frac{2}{n})} = \left(\frac{\theta^{\frac{2}{n}}}{(1+\frac{2}{n})}\right)^n$

$= \frac{\theta^2}{(1+\frac{2}{n})^n}$

$E(x_i^{\frac{2}{n}}) = \int_0^{\theta} x^{\frac{2}{n}} \cdot \frac{1}{\theta} dx$

$= \frac{1}{\theta} \left[\frac{x^{\frac{2}{n}+1}}{\frac{2}{n}+1} \right]_0^{\theta}$

$= \frac{1}{\theta} \left[\frac{\theta^{\frac{2}{n}+1}}{\frac{2}{n}+1} \right]$

$= \frac{\theta^{\frac{2}{n}}}{(1+\frac{2}{n})}$

$$= \overline{\left(1 + \frac{2}{n}\right)^n}$$

$$= \frac{\theta^{2/n}}{\left(1 + \frac{2}{n}\right)}$$

$$\lim_{n \rightarrow \infty} E(T^2) = \lim_{n \rightarrow \infty} \frac{\theta^2}{\underbrace{\left(1 + \frac{2}{n}\right)^n}_{= e^2}} = \frac{\theta^2}{e^2}$$

$$\lim_{n \rightarrow \infty} \left(1 + \frac{1}{n}\right)^n = e$$

$$\lim_{n \rightarrow \infty} \left(1 + \frac{\mu}{n}\right)^n = e^\mu$$

$$\text{check: } \lim_{n \rightarrow \infty} \text{Var}(T) = \lim_{n \rightarrow \infty} E(T^2) - \lim_{n \rightarrow \infty} [E(T)]^2$$

$$= \frac{\theta^2}{e^2} - \frac{\theta^2}{e^2} = 0$$

Q. Let X_1, X_2, \dots, X_n be i.i.d. from $f(x) = e^{-(x-\theta)}$, $x \geq \theta$.

Define $T_n = X_{(1)}$. Show $\lim_{n \rightarrow \infty} E(T_n) = \theta$.

$$X_1 \neq X_{(1)}$$

i.i.d.: X_1, X_2, \dots, X_n [X_i 's can be in any order].

Ordered sample: $X_{(1)} \leq X_{(2)} \leq \dots \leq X_{(n)}$ [Order statistic].

$$\therefore X_{(1)} = \min \{X_1, X_2, \dots, X_n\}$$

$$X_{(n)} = \max \{X_1, X_2, \dots, X_n\}$$

pdf of $X_{(1)}$:

$$T_n = X_{(1)} \xrightarrow{\text{pdf}}$$

$$E(T_n) =$$

$$\text{cdf of } X_{(1)}: P[X_{(1)} \leq y]$$

$$= P[\min\{X_1, X_2, \dots, X_n\} \leq y]$$

$$= 1 - P[\min\{X_1, X_2, \dots, X_n\} \geq y]$$

$$= 1 - P[X_1 \geq y, X_2 \geq y, \dots, X_n \geq y]$$

$$\begin{aligned}
&= 1 - P[X_1 \geq y, X_2 \geq y, \dots, X_n \geq y] \\
&= 1 - P[X_1 \geq y] \cdot P[X_2 \geq y] \dots P[X_n \geq y] \\
&= 1 - \prod P(X_i \geq y) \\
&= 1 - \{P(X_1 \geq y)\}^n \\
&= 1 - \{1 - \underbrace{P(X_1 \leq y)}\}^n \\
&\quad \quad \quad \rightarrow \text{cdf of the popln.}
\end{aligned}$$

$$f(x) = e^{-(x-\theta)}, \quad x \geq \theta.$$

$$P(X_1 \leq y) = \int_{\theta}^y e^{-(x-\theta)} dx.$$

HW. \therefore pdf of $X_{(1)}$ = $\frac{d}{dx} F[X_{(1)}] = n [1 - F(x)]^{n-1} \cdot f(x)$.