

90623  
95723

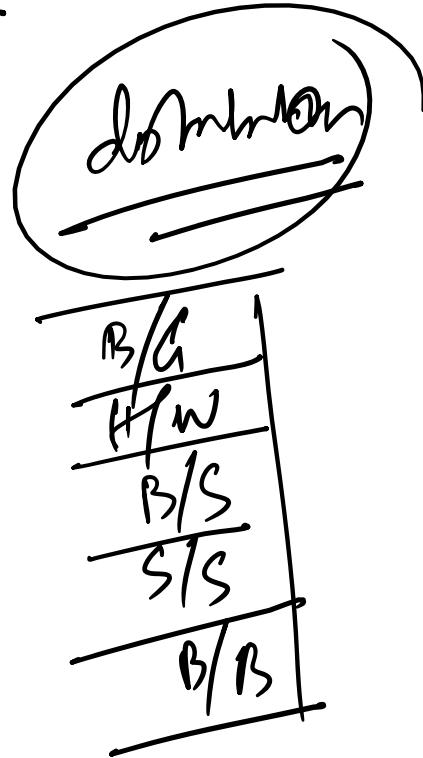
different algorithm in  
Stata8/9

Relationship's  
same

d



$$\begin{aligned}
 & l_{10} > 10 \\
 & l_{10} > l_{100} > l_{10}^{10} \\
 & l_{10}^{10^2} > l_{10}^{10} \\
 & 2l_{10}^{10} > l_{10}^{10} \\
 & 271
 \end{aligned}$$



GAM COM ness

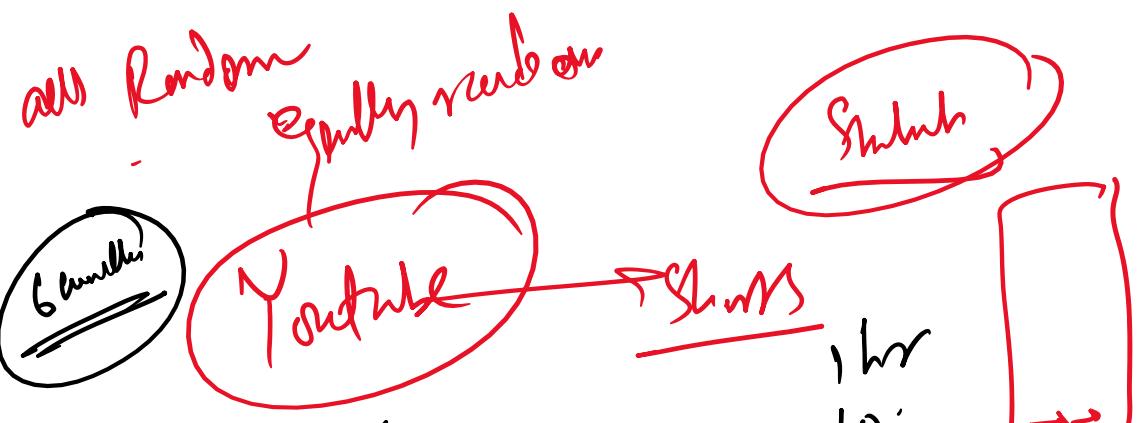
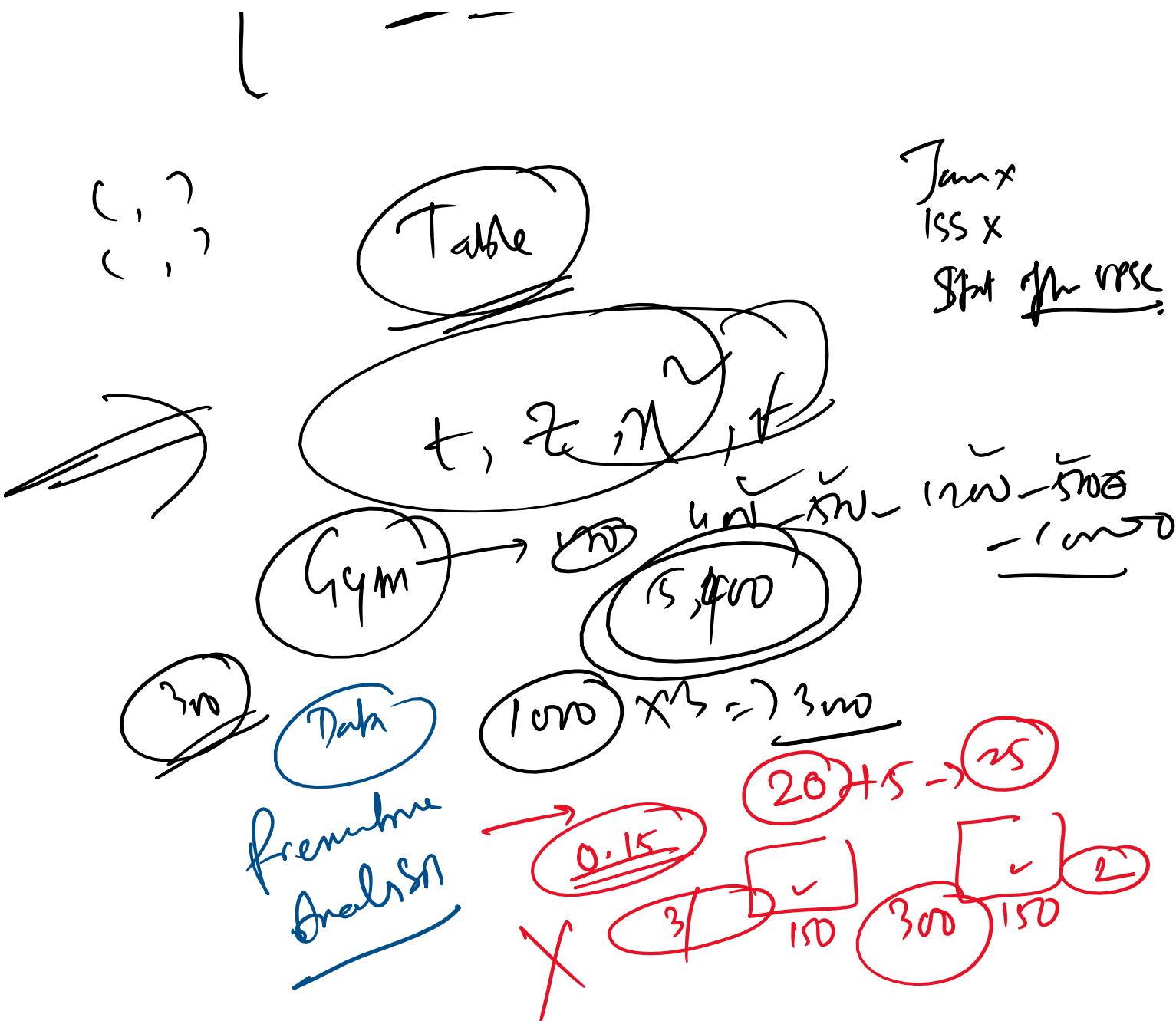
4, 5, 19, 23 -

4, 100, 163, 199, 200 ..

4, -10, 7, 15, -11, ...

$t_1$ ,  $R^{x^2}$   $\rightarrow$   $x \mapsto$

$\ln x, \lg y$



A hand-drawn diagram consisting of three overlapping circles. The leftmost circle contains the number '124'. The middle circle contains the number '126'. The rightmost circle contains the number '120'. Above the circles, there are two separate curved lines. The left curve is black and has two short tick marks on it. The right curve is red and has one short tick mark on it.

 Statistics  
20.7.23 Q...

- Let  $x_1 = 0, x_2 = 1, x_3 = 2, x_4 = 3$  and  $x_5 = 0$  be the observed values of a random sample of size 5 from a discrete distribution with the probability mass function

$$f(x; \theta) = P(X = x) = \begin{cases} \frac{\theta}{3}, & \text{if } x = 0, \\ \frac{2\theta}{3}, & \text{if } x = 1, \\ \frac{1-\theta}{2}, & \text{if } x = 2, 3, \end{cases}$$

where  $\theta \in [0, 1]$  is the unknown parameter. Then the maximum likelihood estimate of  $\theta$  is

- (a)  $\frac{2}{5}$  (b)  $\frac{3}{5}$  (c)  $\frac{5}{6}$  (d)  $\frac{5}{9}$ .

$$\begin{aligned} L(\theta) &= \theta^{x_1} (1-\theta)^{n-x_1} \cdot \theta^{x_2} (1-\theta)^{n-x_2} \cdots \theta^{x_5} (1-\theta)^{n-x_5} = \frac{\theta^3 (1-\theta)^2}{54} \quad \theta \in [0, 1] \\ \ln L(\theta) &= 3 \ln \theta + 2 \ln(1-\theta) - \ln 54 \\ l'(\theta) &= \frac{3}{\theta} - \frac{2}{1-\theta} = 0 \quad \theta = \frac{3}{5} \\ l''(\theta) &= -\frac{3}{\theta^2} - \frac{2}{(1-\theta)^2} < 0 \quad \text{Max} \end{aligned}$$



$$L(\theta) = \prod_{t=1}^T$$

$$f(x, \theta) = \frac{1}{\theta^3} e^{-\sum x_i \theta}$$

$\approx 0$ , when  $x$

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6). Let  $x_1 = 1.1$ ,  $x_2 = 2.2$  and  $x_3 = 3.3$  be the observed values of a random sample of size three from a distribution with the probability density function

$$f(x; \theta) = \begin{cases} \frac{1}{\theta} e^{-x/\theta}, & \text{if } x > 0, \\ 0, & \text{otherwise,} \end{cases}$$

where  $\theta \in \Theta = \{1, 2, \dots\}$  is the unknown parameter. Then the maximum likelihood estimate of  $\theta$  equals

$$P(X_1 = n) = \frac{1}{4} \left(\frac{3}{4}\right)^{n-1}$$

3. Let  $X_1, X_2, X_3$  and  $X_4$  be i.i.d. discrete random variables with the probability mass function

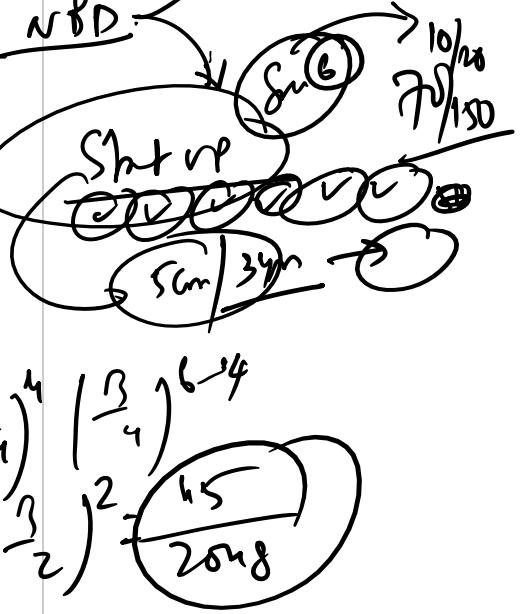
$$P(X_1 = n) = \begin{cases} \frac{3^{n-1}}{4^n}, & \text{if } n = 1, 2, \dots, \\ 0, & \text{otherwise.} \end{cases}$$

Then  $P(X_1 + X_2 + X_3 + X_4 = 6)$  equals \_\_\_\_\_

$$P(Y=6) = {}^4C_{4-1} \frac{\left(\frac{1}{4}\right)^4}{\left(\frac{3}{4}\right)^4} \left(\frac{3}{4}\right)^{4-4}$$

Required Ans..

$$\begin{aligned} P(Y=6) &= {}^6C_{6-1} \left(\frac{1}{4}\right)^4 \left(\frac{3}{4}\right)^{6-4} \\ &= {}^6C_5 \left(\frac{1}{4}\right)^4 \left(\frac{3}{4}\right)^2 \end{aligned}$$



$$E(\max\{X, 5\}) = \sum_{n=1}^{\infty} \{ \max\{n, 5\} \} p(X=n)$$

$$= \sum_{n=1}^{10} (\max\{n, 5\}) \frac{1}{10}$$

Let  $X$  be a random variable with the probability mass function

$$P(X=n) = \begin{cases} \frac{1}{10}, & \text{if } n = 1, 2, \dots, 10, \\ 0, & \text{otherwise.} \end{cases}$$

Then  $E(\max\{X, 5\})$  equals ...

$$\begin{aligned} &= \frac{1}{10} \left( \sum_{n=1}^{10} \max\{n, 5\} \right) + \sum_{n=1}^{10} \max\{n, 5\} \\ &= \frac{1}{10} \left( \sum_{n=1}^{10} 5 + \sum_{n=6}^{10} n \right) \\ &\approx \frac{1}{10} (25 + 60) = 6.5 \end{aligned}$$

$$f(x|\theta) = \frac{1}{\theta^2} e^{-x/\theta} \quad 0 < x < \theta$$

$\in (0, \theta)$

⑤

Let  $X$  be a sample observation from  $U(\theta, \theta^2)$  distribution, where  $\theta \in \Theta = \{2, 3\}$  is the unknown parameter. For testing  $H_0 : \theta = 2$  against  $H_1 : \theta = 3$ , let  $\alpha$  and  $\beta$  be the size and power, respectively, of the test that rejects  $H_0$  if and only if  $X \geq 3.5$ . Then  $\alpha + \beta$  equals \_\_\_\_\_

$$\begin{aligned}\alpha + \beta &= P(X \geq 3.5 | H_0) + P(X \geq 3.5 | H_1) \\ &= P\left(\int_{3.5}^{\infty} f(x|2) dx + \int_{3.5}^{\infty} f(x|3) dx\right) \\ &= \int_{3.5}^{4} \frac{1}{2} dx + \int_{3.5}^{9} \frac{1}{6} dx \\ &= \frac{4 - 3.5}{2} + \frac{9 - 3.5}{2} = \frac{7}{6} = 1.167\end{aligned}$$

Sum of the two lower  $\Rightarrow$  1.167



1.167

$\beta \in 1 - \text{mpfI Err..}$

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Let  $X$  be an observation from a population with density

$$f(x) = \begin{cases} \lambda^2 x e^{-\lambda x}, & \text{if } x > 0, \lambda > 0, \\ 0, & \text{otherwise.} \end{cases}$$

For testing  $H_0 : \lambda = 2$  against  $H_1 : \lambda = 1$ , the most powerful test of size  $\alpha$  is given by "Reject  $H_0$  if  $X > c$ ", where  $c$  is given by

- (a)  $\frac{1}{4} \chi_{4,\alpha}^2$  (b)  $\frac{1}{4} \chi_{5,\alpha}^2$  (c)  $\frac{1}{4} \chi_{2,\alpha}^2$  (d)  $\frac{1}{4} \chi_{1,\alpha}^2$

1 or in 1 H<sub>0</sub> is true)

For testing  $H_0 : \lambda = 2$  against  $H_1 : \lambda = 1$ , the most powerful test of size  $\alpha$  is given by "Reject  $H_0$  if  $X > c$ ".

where  $c$  is given by

- (a)  $\frac{1}{4} \chi_{4,\alpha}^2$  (b)  $\frac{1}{4} \chi_{3,\alpha}^2$  (c)  $\frac{1}{4} \chi_{2,\alpha}^2$  (d)  $\frac{1}{4} \chi_{1,\alpha}^2$ .

$$\text{Size } \alpha = P(\text{Reject } H_0 | H_0 \text{ is true}) \\ = P(X > c | \lambda = 2)$$

$$= \int_c^\infty 4x e^{-2x} dx \\ = \int_c^\infty \frac{1}{\Gamma(4)} y e^{-y} \frac{1}{4} dy \\ = \int_c^\infty \frac{1}{\Gamma(4)} y^3 e^{-y} \frac{1}{4} dy$$

Let  
 $u = y$

$$= P(Y > u)$$

$$u \leq \underline{\chi_{4,\alpha}^2} \quad \text{or, } c = \underline{\chi_{4,\alpha}^2}$$

$$\frac{1}{4} \chi_{4,\alpha}^2$$

②

Ten percent of bolts produced in a factory are defective. They are randomly packed in boxes such that each box contains 3 bolts. Four of these boxes are bought by a customer. The probability that the boxes that this customer bought have no defective bolt in them, is equal to \_\_\_\_\_.

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Suppose that  $F$  is a cumulative distribution function, where

$$F(x) = \begin{cases} 0, & \text{if } x < 0, \\ 1 - e^{-x}, & \text{if } 0 \leq x < 1, \\ c, & \text{if } 1 \leq x < 2, \\ 1 - e^{-x}, & \text{if } x \geq 2. \end{cases}$$

- (i) Find all possible values of  $c$ .
- (ii) Find  $P(0.5 \leq X \leq 2.5)$  and  $P(X = 1) + P(X = 2)$ .

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Let  $X$  be a random variable of continuous type with probability density function

$$f(x|\theta) = \begin{cases} \frac{\theta}{2} \left(\frac{3}{2}\right)^{\theta}, & \text{if } x > 3 \\ 0, & \text{otherwise} \end{cases}; \quad \theta > 0.$$

Based on single observation  $X$ , the most powerful test of size  $\alpha = 0.1$ , for testing  $H_0 : \theta = 1$  against  $H_1 : \theta = 2$ , rejects  $H_0$  if  $X < k$ . Then the value of  $k$  is

- (a) 1    (b)  $\frac{10}{3}$     (c)  $\frac{11}{3}$     (d) 4.

10

Let  $X_1, X_2, \dots$  be a sequence of i.i.d.  $U(0, 1)$  random variables. If  $Y_n = \sum_{i=1}^n X_i$ ,  $n = 1, 2, \dots$ , then

$$\lim_{n \rightarrow \infty} P\left(Y_n \leq \frac{n}{2} + \sqrt{\frac{n}{12}}\right) =$$

- (a) 0.9413   (b) 0.7413   (c) 0.8413   (d) 0.6413.

31. Let  $S_n$  denote the number of heads obtained in  $n$  independent tosses of a fair coin. Using Chebyshev's inequality, the smallest value of  $n$  such that

$$P\left(\left|\frac{S_n}{n} - \frac{1}{2}\right| \leq 0.1\right) \geq \frac{3}{4},$$

is

- (a) 400
- (b) 200
- (c) 300
- (d) 100.

3. Consider the problem of testing  $H_0 : \theta = 0$  against  $H_1 : \theta = 1/2$  based on a single observation  $X$  from  $U(\theta, \theta + 1)$  population. The power of the test "Reject  $H_0$  if  $X > \frac{2}{3}$ " is  
(a)  $\frac{1}{6}$  (b)  $\frac{2}{3}$  (c)  $\frac{1}{3}$  (d)  $\frac{2}{3}$ .

18. Let  $X$  be a single observation from a population having an exponential distribution with mean  $\frac{1}{\lambda}$ . Consider the problem of testing  $H_0 : \lambda = 2$  against  $H_1 : \lambda = 4$ . For the test with rejection region  $X \geq 3$ , let  $\alpha = P(\text{Type I error})$  and  $\beta = P(\text{Type II error})$ . Then
- (a)  $\alpha = e^{-6}$  and  $\beta = 1 - e^{-12}$
  - (b)  $\alpha = e^{-12}$  and  $\beta = 1 - e^{-6}$
  - (c)  $\alpha = 1 - e^{-12}$  and  $\beta = e^{-6}$
  - (d)  $\alpha = e^{-6}$  and  $\beta = e^{-12}$ .

45. A system comprising of  $n$  identical components works if at least one of the components works. Each of the components works with probability 0.8, independent of all other components. The minimum value of  $n$  for which the system works with probability at least 0.97 is \_\_\_\_\_.

53. Let  $X_1, \dots, X_n$  be independent and identically distributed random variables with  $U(0, 1)$  distribution. Then

$$\lim_{n \rightarrow \infty} P\left(\sum_{i=1}^n X_i \leq \frac{n}{2} + n^{3/4}\right) = \text{_____}.$$