

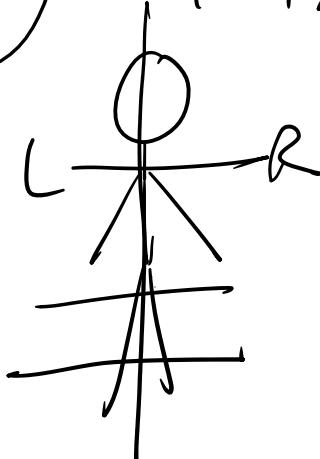
STATISTICAL DISTRIBUTIONS

For binomial process \bigcirc

$H, TH, TTH, \underline{TTTH}, TTTTH, TTTTTH,$

$$\frac{1}{2}, \frac{1}{2} \cdot \frac{1}{2}, \frac{1}{2} \cdot \frac{1}{2}^3, \dots$$

$$\alpha \text{ & } p \rightarrow \left(\frac{\alpha}{1-\alpha} \right) \quad \frac{r_2}{1-r_2} = \frac{r_2}{r_2} = 1$$



Normal symmetry

~~#~~

$$- \frac{(n+1)p-1}{\text{mode}} < r < \frac{(n+1)p}{\text{mode}}$$

90623-95723

Adjusted Poisson..

$$X \sim P(\lambda)$$

$$(1+\kappa)^{-1}$$

Exponential law

$$e^{\lambda} = e^1 = \underline{2 \cdot 718}$$

$$X \rightarrow P(X = \text{Even}) = ?$$

$$p(x) = \frac{e^{-\lambda} \lambda^x}{x!}$$

$$\begin{aligned} P(x = \text{Even}) &= p(0) + p(2) + p(4) + \dots \\ &= e^{-\lambda} \left[1 + \frac{\lambda^2}{2!} + \frac{\lambda^4}{4!} + \dots \right] \end{aligned}$$

Now

$$e^{\lambda} = 1 + \lambda + \frac{\lambda^2}{2!} + \frac{\lambda^3}{3!} + \frac{\lambda^4}{4!} + \dots$$

$$e^1 = 1 + 1 + \underline{(0 \cdot 718)}$$

$$e^{-\lambda} = 1 - \lambda + \frac{\lambda^2}{2!} - \frac{\lambda^3}{3!} + \dots$$

$$e^{\lambda} + e^{-\lambda} = 2 \left[1 + \frac{\lambda^2}{2!} + \frac{\lambda^4}{4!} + \dots \right]$$

$$P(x = \text{Even}) =$$

$$\begin{aligned} &e^{-\lambda} \left[\frac{1}{2} (e^{\lambda} + e^{-\lambda}) \right] \\ &\Rightarrow \underline{\frac{1}{2} (1 + e^{-2\lambda})} \end{aligned}$$

Explains of derivative

(1) P.D

(2) D.E

(3) D.I

(4) ~~B.G.F~~

$X \sim P(\lambda)$

$$M_X(t) = e^{\lambda(e^t - 1)} \quad \lambda > 0.$$

$$E(X) = \lambda$$

Central moment = MGF about Mean λ

$$M_{X-\lambda}(t) = E[e^{t(x-\lambda)}] = e^{-\lambda t} E(e^{tx})$$

$$= e^{-\lambda t} M_X(t)$$

Now, $M_{X-\lambda}(t) = e^{-\lambda t} \cdot e^{\lambda(e^t - 1)} = e^{\lambda(e^t - 1 - t)}$

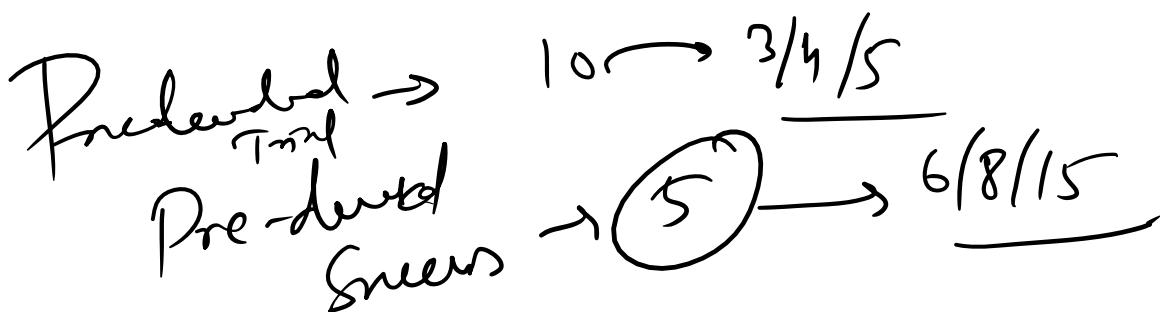
$$M(t) = e^{\lambda(e^t - 1 - t)}$$

$$\frac{dM(t)}{dt} = e^{\lambda(e^t - 1 - t)} \cdot \lambda(e^t - 1)$$

$$\frac{d}{d\lambda} M(t) = e^{\lambda(e^t - 1 - t)} (e^t - 1 - t)$$

$$= M(t) (e^t - 1 - t)$$

~~H~~ ~~E~~



Ques ... $x, y \sim \text{ind}$ $\xrightarrow{\text{R.V.}}$ $x+y \sim P(\lambda+\mu)$

$X \sim P(\lambda)$, find the distribution of Y .

Ans: $X \sim P(\lambda)$, $X+Y \sim P(\lambda+\mu)$

$$M_X(t) = e^{\lambda(e^{t-1}-1)} \quad \text{--- (1)}$$

$$M_{X+Y}(t) = e^{(\lambda+\mu)(e^{t-1}-1)} \quad \text{--- (2)}$$

$$\text{So, } M_{X+Y}(t) = M_X(t) \cdot M_Y(t)$$

$$\Rightarrow e^{(\lambda+\mu)(e^{t-1}-1)} = \frac{e^{\lambda(e^{t-1}-1)}}{e^{\lambda(e^{t-1}-1)}} M_Y(t) = e^{\mu(e^{t-1}-1)}$$

Ques Show that, X is a Poisson variate mean λ

$$E(X^2) = \lambda + (X+1) \neq$$

$$E(X^2) = \text{var}(X) + [E(X)]^2 = \lambda + \lambda^2$$

$$\lambda(1+\lambda) = \lambda[1+E(X)]$$

$$E(X^2) = \lambda [E(X)+1] = \lambda E(X+1)$$

$$\text{If } \lambda=1, E|X-1| = \text{MD of PD about mean}$$

$$= \sum_{x=0}^{\infty} |x-1| p(x)$$

$$\approx \sum_{x=0}^{\infty} |x-1|$$

$$\begin{aligned}
 &= \sum_{x=0}^{\infty} \left(\frac{e^{-1} x}{x!} \right) \stackrel{x \geq 0}{\overbrace{|x-1|}} = e^{-1} \sum_{x=0}^{\infty} \frac{|x-1|}{x!} \\
 &= e^{-1} \left[1 + 0 + \frac{1}{2!} + \frac{2}{3!} + \frac{3}{4!} + \dots \right] \\
 &= e^{-1} \left[1 + \frac{1}{2!} + \frac{2}{3!} + \frac{3}{4!} + \dots \right] \\
 \cancel{\text{Now}} \\
 \frac{n}{(n+1)!} &= \frac{(n+1)-1}{(n+1)!} = \frac{(n+1)}{(n+1)!} - \frac{1}{(n+1)!} \\
 &= \frac{1}{n!} - \frac{1}{(n+1)!} \\
 \therefore E|x-1| &= e^{-1} \left[\left(1 + \left(1 - \frac{1}{2!} \right) \right) + \left(\frac{1}{2!} - \frac{1}{3!} \right) + \dots \right] \\
 &= e^{-1} \left[(1+1) \right] = \underline{\frac{2}{e} \cdot \text{mean}}
 \end{aligned}$$

More Many times + ~~2~~
Random & Partition Structure
Negative Binomial Binomials + Geometric Sub
10 → 4 10C4 p⁶ q⁴

Note on Q/H ..
Improper Integral .. → 3 types
x. $\int_{r/1}^{\infty}$ Dom (0, ∞)

$$\int_0^\infty e^{-x} dx \stackrel{\text{Improper}}{\int} + \text{dom } [0, \infty)$$

(I) When domain > finite

$$\int_0^\infty \frac{1}{x^{n-1}} dx - \int_1^\infty \frac{1}{(x-1)^n}$$

finite value \rightarrow finite Ans.

$$\int_1^\infty \frac{dx}{x-1} \quad \int_1^\infty \frac{dx}{(x-1)^n} \quad \text{Cont/Disc}$$

(II) Mmp of I + P

$$\int_0^\infty \frac{dx}{x(1-x)}$$

$$B(m, n) = \int_0^1 x^{m-1} (-x)^{n-1} dx$$

Cont

$$B(2, 2) = \int_0^1 x^{2-1} (1-x)^{2-1} dx$$

$$B(m, n) = \frac{(m-1)!}{n(n+1)\dots(n+m-1)}$$

$$B(m, n) = \int_0^\infty \frac{x^{m-1}}{(1+x)^{m+n}} dx$$

Gamma function:

the integral

$\Gamma(n) = \int_0^{\infty} e^{-x} x^{n-1} dx$

Cont. n > 0

$$\Gamma(n+1) = \Gamma(n)$$

$$\beta(m, n) = \frac{\Gamma(m)\Gamma(n)}{\Gamma(m+n)}$$

$$\Gamma\left(\frac{1}{2}\right) = \sqrt{\pi}$$

$$\Gamma(1-m)\Gamma(m) =$$

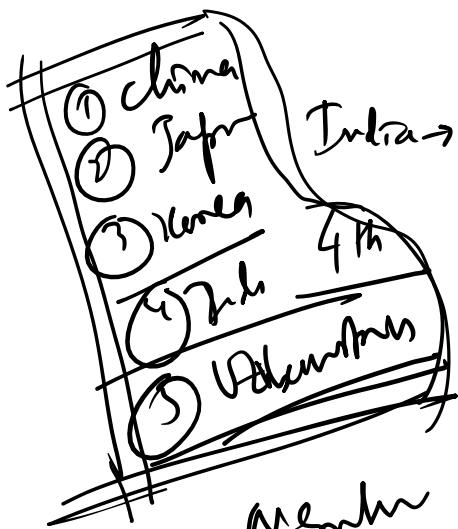
$$\int_0^{\infty} e^{-x^2} x^n dx = ?$$

But $x^2 = z$, $2x dx = dz$ $dx = \frac{dz}{2}$

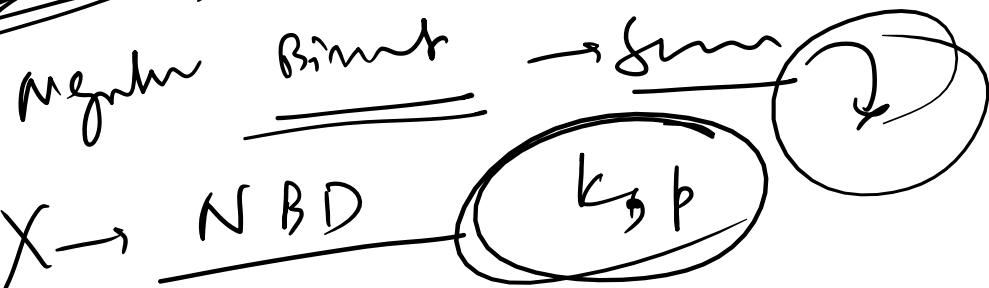
at $x=0$, $z=0$

$n=\infty$, $z=\infty$

$$\begin{aligned} \therefore \int_0^{\infty} e^{-x^2} x^n dx &= \int_0^{\infty} e^{-z} \cdot z^{\frac{n}{2}} \cdot \frac{1}{2} z^{\frac{1}{2}} dz \\ &= \frac{1}{2} \int_0^{\infty} e^{-z} \cdot z^{\frac{n}{2}} dz = \frac{1}{2} \int_0^{\infty} e^{-z} z^{5-\frac{1}{2}} dz \\ &\geq \frac{1}{2} \Gamma(5) = \frac{1}{2} \times 4! = 12 \end{aligned}$$



India is outcome $\rightarrow \textcircled{1} \rightarrow 3/4/5/6/7/8\ldots$



$$P(X \geq m) = ?$$

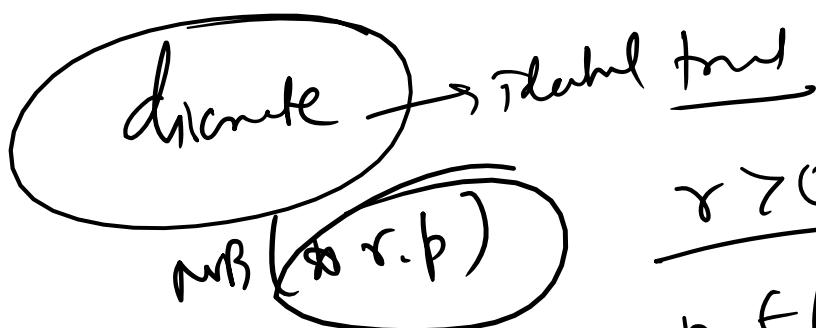
$X \sim NB(k, p) \rightarrow pmf \rightarrow$

$$f(x) = \frac{-k}{n} C_n^x p^n (-q)^{n-x}$$

$n = 0, 1, 2, \dots$

$$\text{where, } P = \frac{p}{q}, \quad q = \frac{1-p}{p} \quad p+q=1 \quad q-p=1$$

$n = 0 \text{ & } N$



$r \geq 0$ no of some initial or endpoint is always

$$p \in (0, 1)$$

$$k = \text{number } \in \{0, 1, 2, 3, \dots\}$$

\dots, r

$$K = \text{number of } \{0, 1, 2, \dots\} \\ \text{Pmf} \rightarrow {}_{K+r-1}C_K \cdot (1-p)^K p^r$$

$$\text{Mean} \rightarrow \frac{r(1-p)}{p}$$

$$\text{Var} \rightarrow \frac{r(1-p)}{p \cdot p}$$

~~Moment Methods~~

$$\sigma = \sqrt{\frac{E[X]^2}{V[X]} - E[X]}$$

$$p = 1 - \frac{E[X]}{\sqrt{V[X]}}$$



Related to the concept of
working for a fixed number of
trials in Bernoulli Trials

→ failure before the succ B ..
(~~if~~ -ve) → Counting of failures